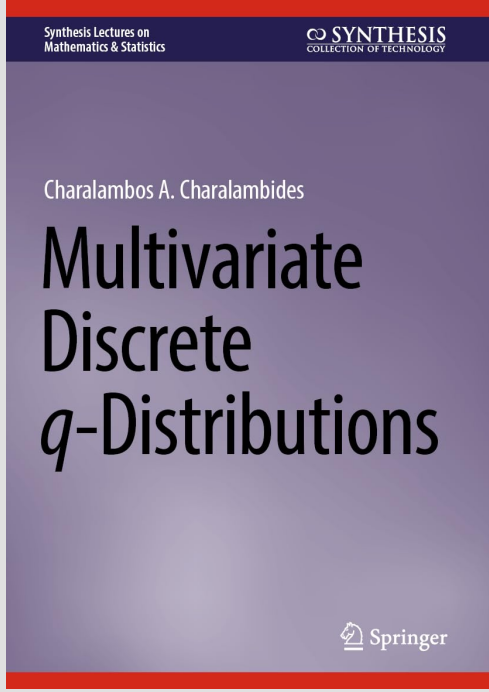
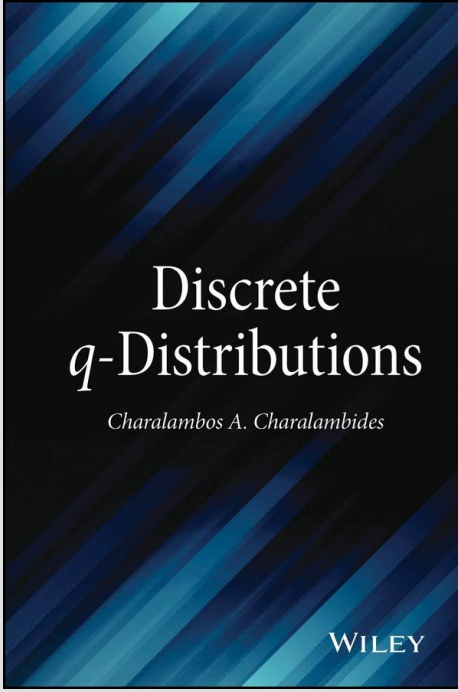
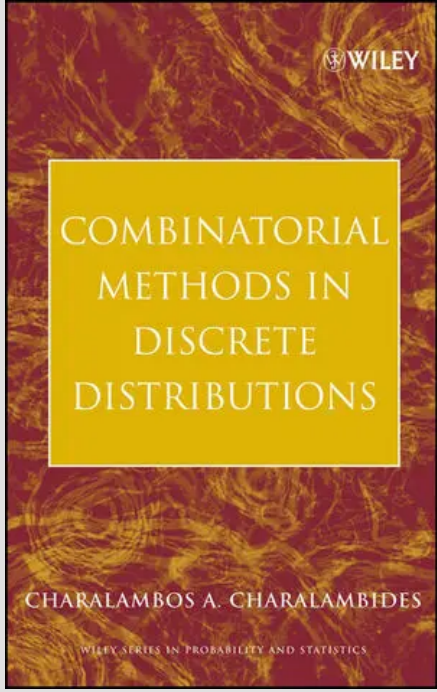
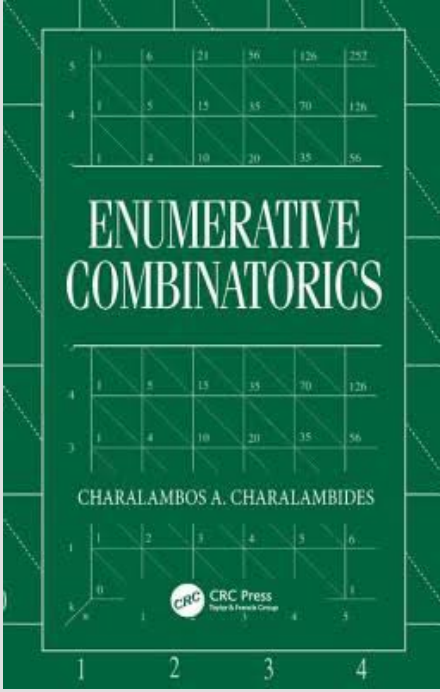


Charalambos A. Charalambides (1945–2024)

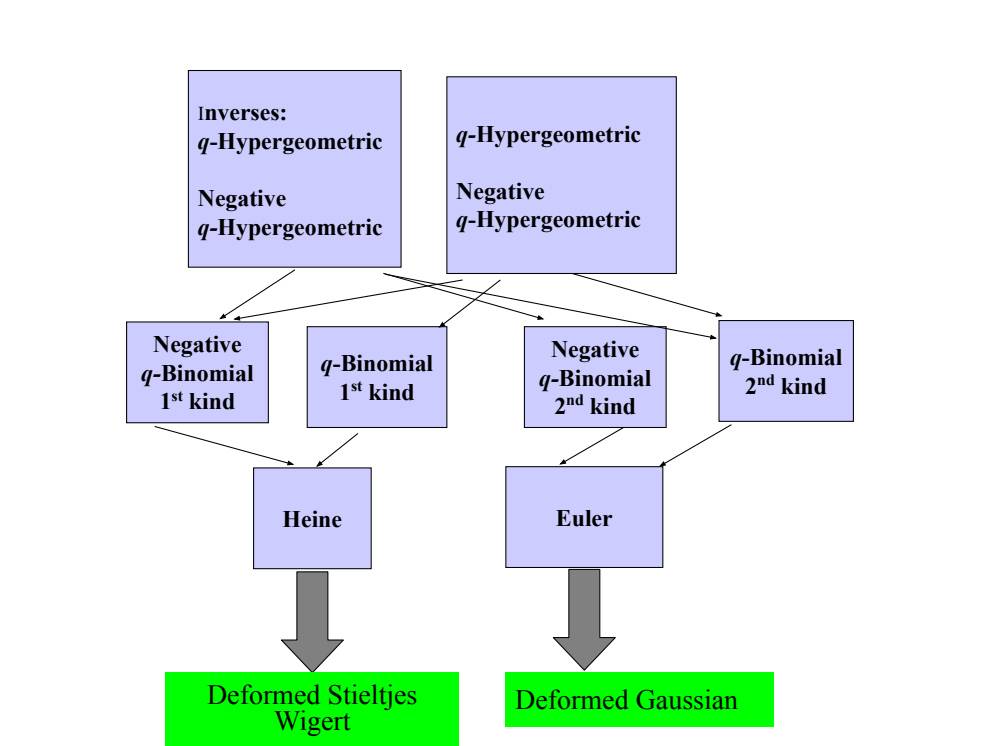


- Born: 19 December 1945, Koilani, Limassol, Cyprus.
- B.Sc. in Mathematics, National and Kapodistrian University of Athens (1969).
- Ph.D. in Statistics (1972), supervised by Professor Th. Kakoullos.
- Joined the Department of Mathematics in 1971; full professor in 1987.
- Visiting professor at McGill University, Temple University, and the University of Cyprus.
- Research interests: [combinatorics](#), [discrete distributions](#), [parameter estimation](#), [probability and statistics](#).
- Author of four research monographs, several textbooks, and ~ 70 journal articles.
- His scientific legacy continues to inspire researchers and students.

Research monographs in enumerative combinatorics and q-distributions

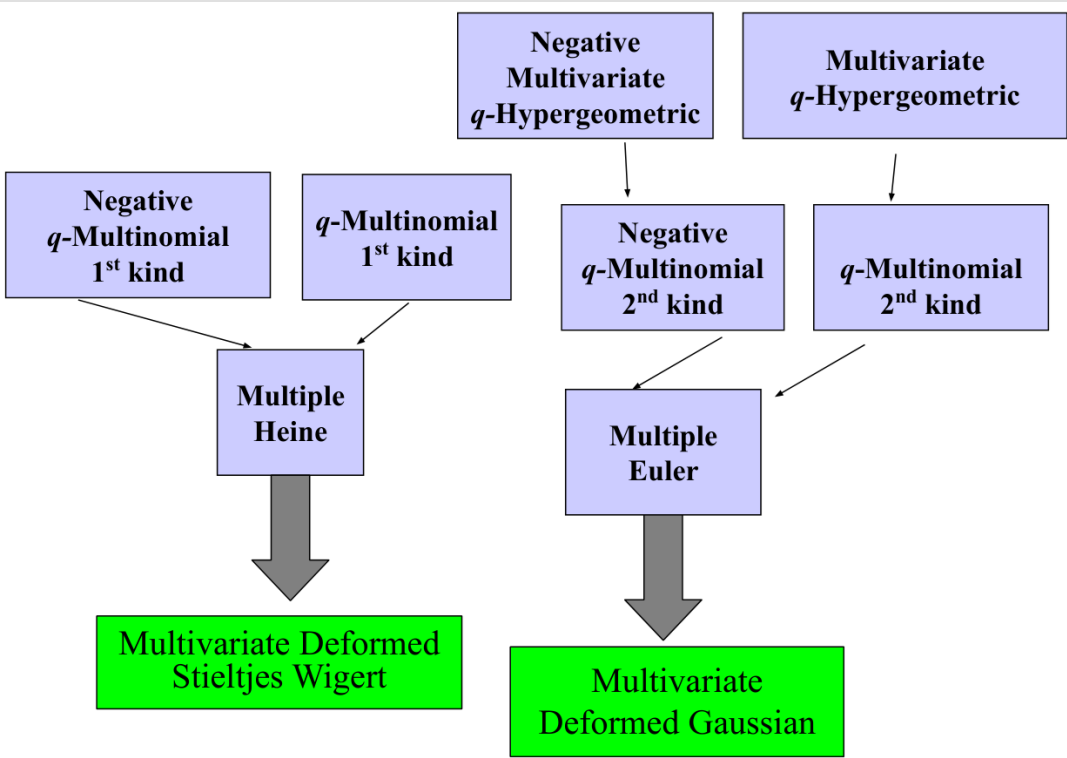


Univariate and multivariate discrete q-distributions (Charalambides, 2016, 2024)



Univariate discrete *q*-distributions:

stochastic models of sequences of n independent Bernoulli trials with success probability varying geometrically (with rate q) either with the number of previous trials or with the number of previous successes, or both.



Multivariate discrete *q*-distributions:

stochastic models of sequences of n independent Bernoulli trials with chain-composite successes, where the odds of success of a certain kind at a trial is assumed to vary geometrically (with rate q) with the number of previous trials or with the number of previous successes, or both.

q-Multinomial distribution of the 1st kind

- X_j : #successes of a j th kind in a sequence of n independent Bernoulli trials with chain composite failures, where the probability of success of the j th kind at the i th trial is given by

$$p_{j,i} = \frac{\theta_j q^{i-1}}{1 + \theta_j q^{i-1}}, \quad 0 < \theta_j < \infty, \quad j \geq 1, i \geq 1, \quad 0 < q < 1.$$

- Joint probability function of the random vector $\mathcal{X} = (X_1, \dots, X_k)$ is given by

$$f_{\mathcal{X}}^B(x_1, \dots, x_k) = P(X_1 = x_1, \dots, X_k = x_k) = \binom{n}{x_1, \dots, x_k}_q \prod_{j=1}^k \frac{\theta_j^{x_j} q^{\binom{x_j}{2}}}{\prod_{i=1}^{n-s_{j-1}} (1 + \theta_j q^{i-1})}$$

with $\sum_{j=1}^k x_j \leq n$ and $s_j = \sum_{i=1}^j x_i$.

The discrete limit of the joint p.f. of the q -multinomial distribution of the 1st kind is the [multiple Heine distribution](#):

$$\lim_{n \rightarrow \infty} \binom{n}{x_1, \dots, x_k}_q \prod_{j=1}^k \frac{\theta_j^{x_j} q^{\binom{x_j}{2}}}{\prod_{i=1}^{n-s_{j-1}} (1 + \theta_j q^{i-1})} = \prod_{j=1}^k \frac{q^{\binom{x_j}{2}} \lambda_j^{x_j}}{[x_j]_q!} \prod_{i=1}^{\infty} (1 + \lambda_j (1 - q) q^{i-1})^{-1}, \quad x_j = 0, \dots, \quad \lambda_j > 0, \quad \lambda_j = \theta_j / (1 - q).$$

q-Multinomial distribution of the 2nd kind

- Y_j : number of successes of a j th kind in a sequence of n independent Bernoulli trials with chain composite failures, where the conditional probability of success of the j th kind at any trial, given that $i - 1$ successes of the j th kind occur in the previous trials is given by

$$p_{j,i} = 1 - \theta_j q^{i-1}, \quad j \geq 1, i \geq 1, \quad 0 < \theta_j < 1.$$

- Joint probability function of the random vector $\mathcal{Y} = (Y_1, \dots, Y_k)$:

$$f_{\mathcal{Y}}^{MS}(y_1, \dots, y_k) = P(Y_1 = y_1, \dots, Y_k = y_k) = \binom{n}{y_1, \dots, y_k}_q \prod_{j=1}^k \theta_j^{y_j} \prod_{i=1}^{n-s_j} (1 - \theta_j q^{i-1}),$$

$y_j = 0, 1, \dots, n, \quad \sum_{j=1}^k x_j \leq n, \quad s_j = \sum_{i=1}^j y_i$.

The discrete limit of the joint p.f. of the q -multinomial distribution of the 2nd kind is the [multiple Euler distribution](#):

$$\lim_{n \rightarrow \infty} \binom{n}{y_1, \dots, y_k}_q \prod_{j=1}^k \theta_j^{y_j} \prod_{i=1}^{n-s_j} (1 - \theta_j q^{i-1}) = \prod_{j=1}^k (1 - \lambda_j (1 - q) q^{i-1}) \frac{\lambda_j^{y_j}}{[y_j]_q!}, \quad \lambda_j = \theta_j / (1 - q).$$

Multivariate *q*-hypergeometric distribution

- Consider an urn containing ν balls, $\{b_1, \dots, b_\nu\}$, of $k + 1$ different ordered colors, with ν_ℓ distinct balls of color c_ℓ , $\{b_{s_{\ell-1}+1}, b_{s_{\ell-1}+2}, \dots, b_{s_\ell}\}$, for $\ell = 1, \dots, k + 1$, where $s_0 = 0$, and

$$s_\ell = \sum_{i=1}^{\ell} \nu_i, \quad \text{for } \ell = 1, \dots, k + 1,$$

with $s_{k+1} = \nu$. The first k colors, $\{c_1, \dots, c_k\}$, may be considered as shades of white, and color $\{c_{k+1}\}$ as black.

- W_ℓ : number of balls of color c_ℓ drawn in n q -drawings in a multiple q -hypergeometric urn model, with the conditional probability of drawing a ball of color c_ℓ at the i th q -drawing, given that $j_\ell - 1$ white balls of color c_ℓ and a total of colors $c_1, \dots, c_{\ell-1}$ are drawn in the previous $i - 1$ q -drawings:

$$p_{i,j_\ell}(\ell) = \frac{[\nu_\ell - j_\ell + 1]_q}{[\nu + s_\ell - i + 1]_q}, \quad j_\ell = 1, \dots, i, \quad \ell = 1, \dots, k, \quad i \geq 1.$$

- Joint probability function of a random vector $\mathcal{W} = (W_1, \dots, W_k)$:

$$f_{\mathcal{W}}(w_1, \dots, w_k) = P(W_1 = w_1, \dots, W_k = w_k) = \binom{n}{w_1, \dots, w_k}_q \frac{q^{\sum_{j=1}^k (n - \sum_{i=1}^j w_i)(w_j - w_j)} \prod_{j=1}^{k+1} [\nu_j]_{w_j, q}}{[\nu]_{n, q}}$$

$w_j = 0, \dots, n, \quad j = 0, \dots, k$, with $\sum_{j=1}^k n_j \leq n$, where $w_k = n - \sum_{j=1}^k w_j, \nu = \sum_{j=1}^{k+1} \nu_j$, and $n \leq [\nu_j], j = 1, \dots, k$.

On *q*-continuous random variables, K & V' 17, V'23

A k -variate random variable $\mathcal{X} = (X_{1,q}, \dots, X_{k,q})$ is called [q-continuous](#) if there exists a non-negative function

$$f_q(x_1, \dots, x_k) \geq 0, \quad x_i > 0, \quad i = 1, \dots, k$$

with

$$\int_0^\infty \cdots \int_0^\infty \int_0^\infty f_q(x_1, \dots, x_k) d_q x_1 \cdots d_q x_k = 1,$$

where for every real numbers α_i, β_i with $\alpha_i < \beta_i, \quad i = 1, \dots, k$

$$P(\alpha_1 < X_{1,q} \leq \beta_1, \dots, \alpha_k < X_{k,q} \leq \beta_k) = \int_{\alpha_k}^{\beta_k} \cdots \int_{\alpha_1}^{\beta_1} f_q(x_1, \dots, x_k) d_q x_1 \cdots d_q x_k.$$

The function $f_q(x_1, \dots, x_k) \geq 0$ is called [q-density function](#) of the k -variate random variable $\mathcal{X} = (X_{1,q}, \dots, X_{k,q})$ or **joint *q*-density function** of the random variables $X_{1,q}, \dots, X_{k,q}$.

For the corresponding joint distribution function, we have:

$$F(x_1, \dots, x_k) = P(X_{1,q} \leq x_1, \dots, X_{k,q} \leq x_k) = \int_0^{x_k} \cdots \int_0^{x_1} f_q(t_1, \dots, t_k) d_q t_1 \cdots d_q t_k.$$

Taking the q -derivative of the above relation we have

$$\frac{\partial F(x_1, \dots, x_k)}{\partial_q x_k \cdots \partial_q x_1} = f_q(x_1, \dots, x_k) = \frac{P(qx_1 < X_{1,q} \leq x_1, \dots, qx_k < X_{k,q} \leq x_k)}{(1 - q)x_1 \cdots (1 - q)x_k}.$$

On Heine stochastic process, K & V' 17

- Consider a stochastic model that is developing in time or space, and let $X(t), t \geq 0$, be the number of successes (occurrence of event A) in the interval $(0, t]$. Assume that $X(t), t \geq 0$, is a stochastic process, with independent and homogeneous increments, which starts at time $t = 0$ from state 0, $P(X(0) = 0) = 1$, and, in the small time interval $(qt, t]$, of length $\delta t = (1 - q)t$, satisfies the condition

$$p_j(\delta t) = P(X(t) - X(qt) = j) = \begin{cases} \frac{1}{1 + \lambda(1 - q)t}, & j = 0, \\ \frac{\lambda(1 - q)t}{1 + \lambda(1 - q)t}, & j = 1, \\ 0, & j > 1, \end{cases}$$

with $0 < \lambda < \infty$. Then, $X(t), t \geq 0$, is called [Heine process](#), with parameters λ and q .

- The probability function of the Heine process, $X(t), t \geq 0$, with parameters λ and q , is given by

$$P(X(t) = k) = e_q(-\lambda t) \frac{q^{\binom{k}{2}} (\lambda t)^k}{[k]_q!},$$

for $k \in \mathbb{N}$, where $e_q(z) = \prod_{i=1}^\infty (1 - (1 - q)zq^{i-1})^{-1}$, $|z| < 1/(1 - q)$ is the small q -exponential function.

- The distribution function of the waiting times $W_n, n \geq 1$, in a Heine process is that of the [q-Erlang distribution of the first kind](#), with parameters n, λ and q , given by

$$F_n(w) = P(W_n \leq w) = 1 - \sum_{x=0}^{n-1} e_q(-\lambda w) \frac{q^{\binom{x}{2}} (\lambda w)^x}{[x]_q!}, \quad 0 < w < \infty,$$

where n is a positive integer, $0 < \lambda < \infty$. Its q -density function $f_n(w) = d_q F_n(w)/d_q w$ is given by

$$f_n(w) = \frac{q^{\binom{n}{2}} \lambda^n}{[n - 1]_q!} t^{n-1} e_q(-\lambda w), \quad 0 < w < \infty.$$

- The q -density function of the interarrival times $T_1 = W_1, T_k = W_k - W_{k-1}, k \geq 2$, in a Heine process is that of the [q-exponential distribution of the first kind](#), given by

$$f_{T_k}(t) = \lambda e_q(-\lambda t), \quad \text{for } 0 < t < \infty \text{ and } 0 < \lambda < \infty.$$

- Let $\mathcal{W} = (W_{1,q}, \dots, W_{\nu,q})$, be a ν -th-variate random variable and $W_{k,q}, k = 1, \dots, \nu$ be the random variables of the waiting times in a Heine process. Then the joint q -density function of the random variables $W_{k,q}, k = 1, \dots, \nu$, is given by

$$f_{\mathcal{W}}(w_1, \dots, w_\nu) = \lambda^\nu \prod_{j=1}^\nu e_q(-\lambda(w_j - w_{j-1})), \quad w_0 = 0,$$

$0 < w_1 < \dots < w_\nu < \infty$, where $0 < \lambda < \infty$.

Motivation: a theorem leading to *q*-order statistics, V'23

Let W_n be the waiting time of the n th arrival of the Heine process $\{X(t), t > 0\}$ with parameters λ and q . Then the joint q -density function of the waiting times $W_1, \dots, W_\nu$, in which the first ν events occur given that $X(t) = \nu, \quad 0 < w_1 < \dots < w_\nu < t$ with $w_i \in (q^{\nu-i+1}t, q^{\nu-i}t], i = 1, \dots, \nu - 1$, is given by

$$f_q(w_1, \dots, w_\nu | X(t) = \nu) = \frac{[\nu]_q!}{q^{\binom{\nu}{2}} t^\nu}.$$

Classical order statistics

Let a ν -variate random variable $\mathcal{X} = (X_1, \dots, X_\nu)$ be defined in a sample space Ω . Then for the values $x_1 = X_1(\omega), \dots, x_\nu = X_\nu(\omega), \omega \in \Omega$ there is a permutation $(i_1, \dots, i_\nu)$ of the ν indices $\{1, \dots, \nu\}$, such that $x_{i_1} \leq \dots < x_{i_{\nu-1}} \leq x_{i_\nu}$. The k -th ordered random variable is denoted by $X_{(k)}$ and defined by $X_{(k)}(\omega) = x_{(k)}, \omega \in \Omega$, where $x_{(k)} = x_{i_k}, k = 1, \dots, \nu$. In particular, for $k = 1$ this gives $X_{(1)} = \min\{X_1, \dots, X_\nu\}$ and, for $k = \nu$ this gives $X_{(\nu)} = \max\{X_1, \dots, X_\nu\}$. Generally, $X_{(1)} \leq \dots \leq X_{(\nu)}$.

q-Order statistics (V'23), dependencies & patterns (V'25)

Let $\mathcal{Y} = (Y_1, \dots, Y_\nu)$ be a ν -variate random variable and $Y_{(k)}$, be the corresponding k -th ordered random variables. Assume

$$0 \leq Y_{(1)} < qY_{(2)} < Y_{(2)} < \dots < Y_{(\nu-1)} < qY_{(\nu)} < Y_{(\nu)} \leq \beta,$$

then, $Y_{(k)}$ is called the k -th *q-ordered* random variables.

For any integer r between 1 and ν , let $\{i_1, \dots, i_r\}$ be an r -combination of $\{1, \dots, \nu\}$ satisfying $i_1 < \dots < i_r$, and let $\{i_{r+1}, \dots, i_\nu\}$ be its complementary combination (i.e., one has $\{i_1, \dots, i_r\} \cup \{i_{r+1}, \dots, i_\nu\} = \{1, \dots, \nu\}$) satisfying $i_{r+1} < \dots < i_\nu$. Then, the non-ordered random variables Y_i 's satisfy the following dependence relations for $y \in [0, \beta]$:

$$P(Y_{i_r} \leq y | Y_{i_1} \leq y, \dots, Y_{i_{r-1}} \leq y) = P(Y_{i_r} \leq q^{\text{des}(\mathcal{Y}_r)} y), \quad \text{where } \text{des}(\mathcal{Y}_r) = \sum_{j=1}^{r-1} \mathbf{1}_{Y_{i_j} > Y_{i_{j+1}}} \text{ with } \mathcal{Y}_r = (Y_{i_1}, \dots, Y_{i_r}),$$

$$P(Y_{i_r} \leq y | Y_{i_1} > y, \dots, Y_{i_{r-1}} > y) = P(Y_{i_r} \leq y), \quad \text{and}$$

$$P(Y_{i_m} \leq y | Y_{i_1} \leq y, \dots, Y_{i_r} \leq y, Y_{i_{r+1}} > y, \dots, Y_{i_{m-1}} > y) = P(Y_{i_m, q} \leq q^{\text{des}(\mathcal{Y}_m)} y), \quad \text{with } \mathcal{Y}_m = (Y_{i_1}, \dots, Y_{i_m}), m = r + 1, \dots, \nu.$$

Results

- Distributional properties of q -order statistics.
- Joint and marginal q -densities for dependent random variables.
- Results for extrema, spacings, ratios, differences, and sample medians.
- Distributional properties for dependent, non-identically distributed q -uniform and discrete q -uniform random variables.
- Permutation statistics and their q -analogues as fundamental tools in the analysis.
- Extensions to q -Dirichlet, q -Beta, and Heine-process models.

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